

Lecture 17:
Finish Review of EXAM I Chapter 16, 1-5

6 weeks left...12 lectures.

I will cover **at least** chpt 6-11. Any spare time will be used in the lab.

Lecture on Chapter 6

Specification:

1. Choosing the correct independent variables
2. choosing the correct functional form
3. Choosing the correct for of the error.

Specification error occurs when an error occurs in the three steps above.

Omitted variables

True regression

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \varepsilon_i$$

estimated

$$Y_i = \beta_0^* + \beta_1^* X_{1i} + \varepsilon_i^*$$

where

$$\varepsilon_i^* = \varepsilon_i + \beta_2 X_{2i}$$

so

$$E(\hat{\beta}_0^*) \neq \beta_0$$

and

$$\text{if } r_{12} = r_{21} = \frac{\sum (x_1 - \bar{x}_1)(x_2 - \bar{x}_2)}{\sqrt{\sum (x_1 - \bar{x}_1)^2 \sum (x_2 - \bar{x}_2)^2}} \neq 0$$

then

$$E(\hat{\beta}_1^*) \neq \beta_1$$

Solution and identification?

Irrelevant variables

True regression

$$Y_i = \beta_0 + \beta_1 X_{1i} + \varepsilon_i$$

estimated

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \varepsilon_i^{**}$$

where

$$\varepsilon_i^{**} = \varepsilon_i - \beta_2 X_{2i}$$

Four Important Specification Criteria

1. Theory
2. T-test
3. Rbar squared
4. Bias (do variables coefficients change significantly when variables are added)

Specification Searches

Data Mining

http://www.absoluteastronomy.com/topics/Testing_hypotheses_suggested_by_the_data

Stepwise regressions

<http://www.stata.com/support/faqs/stat/stepwise.html>

Sequential searches

Using T-tests to choose included variables

Scanning and Sensitivity analysis

So how do we choose a model?

Lecture 18: October 31

Lagged independent variables

Ramsey Regression Specification Error Test (RESET)

A test for misspecification and sometimes, rather mistakenly referred to as a test for omitted variables

Using OLS estimate

$$\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_{1i} + \hat{\beta}_2 X_{2i} \quad \text{eq 1}$$

then generate $\hat{Y}_i^2, \hat{Y}_i^3, \hat{Y}_i^4$

re-estimate the original equation augmenting it with the polynomials of the fitted values.

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 \hat{Y}_i^2 + \beta_4 \hat{Y}_i^3 + \beta_5 \hat{Y}_i^4 + \varepsilon_i \quad \text{eq 2}$$

$$F = \frac{(RSS_m - RSS)/M}{RSS/(n - (k + 1))}$$

where RSS_m is from eq 1 and RSS is from eq 2.

Ramsey's Regression Specification Error Test (RESET)

<http://faculty.chass.ncsu.edu/garson/PA765/assumpt.htm>

□ Ramsey's RESET test (regression specification error test). Ramsey's general test of specification error of functional form is an F test of differences of R^2 under linear versus nonlinear assumptions. It is commonly used in time series analysis to test whether power transforms need to be added to the model. For a linear model which is properly specified in functional form, nonlinear transforms of the fitted values should not be useful in predicting the dependent variable. While STATA and some packages label the RESET test as a test to see if there are "no omitted variables," it is a linearity test, not a general specification test. It tests if any nonlinear transforms of the specified independent variables have been omitted. It does not test whether other relevant linear or nonlinear variables have been omitted.

1. Run the regression to obtain R_o^2 , the original multiple correlation.
2. Save the predicted values (Y 's).
3. Re-run the regression using power functions of the predicted values (ex., their squares and cubes) as additional independents for the Ramsey RESET test of functional form where testing that none of the independents is nonlinearly related to the dependent. Alternatively, re-run the regression using power functions of the independent variables to test them individually.
4. Obtain R_n^2 , the new multiple correlation.
5. Apply the F test, where $F = (R_n^2 - R_o^2)/[(1 - R_n^2)/(n-p)]$, where n is sample size and p is the number of parameters in the new model.
6. Interpret F : For an adequately specified model, F should be non-significant.

Apparently some stats programs have rounding errors/computational problems that appear as multicollinearity. <http://en.wikipedia.org/wiki/Multicollinearity>

4) Mean-center the predictor variables. Mathematically this has no effect on the results from a regression. However, it can be useful in overcoming problems arising from rounding and other computational steps if a carefully designed computer program is not used.

But really, it shouldn't truly matter. <http://www.bauer.uh.edu/jhess/papers/JMRMeanCenterPaper.pdf>

But now that I do some digging I see that stata actually does this normalization as well, before taking the powers.

<http://www.stata.com/statalist/archive/2004-06/msg00264.html>

Akaike Information Criterion (AIC)

Minimize $AIC = \text{Log}(RSS/n) + 2(K+1)/n$

Schwarz Criterion, or Schwarz Bayesian Criterion (SC, SBC)

Minimize $SBC = \text{Log}(RSS/n) + \text{Log}(n)(K+1)/n$

Lecture 19: November 4

The use and interpretation of the constant term

Don't do it. There is an inherent identification problem, as the constant includes the true constant, means of omitted variables, and

Alternative functional forms

Linear Form

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \varepsilon_i$$

Double log form

$$\ln Y_i = \beta_0 + \beta_1 \ln X_{1i} + \beta_2 \ln X_{2i} + \varepsilon_i$$

Semi-log form

Log – Lin

$$\ln Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \varepsilon_i$$

Lin-Log

$$Y_i = \beta_0 + \beta_1 \ln X_{1i} + \beta_2 \ln X_{2i} + \varepsilon_i$$

Polynomial functional form

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{1i}^2 + \varepsilon_i$$

Inverse functional Form

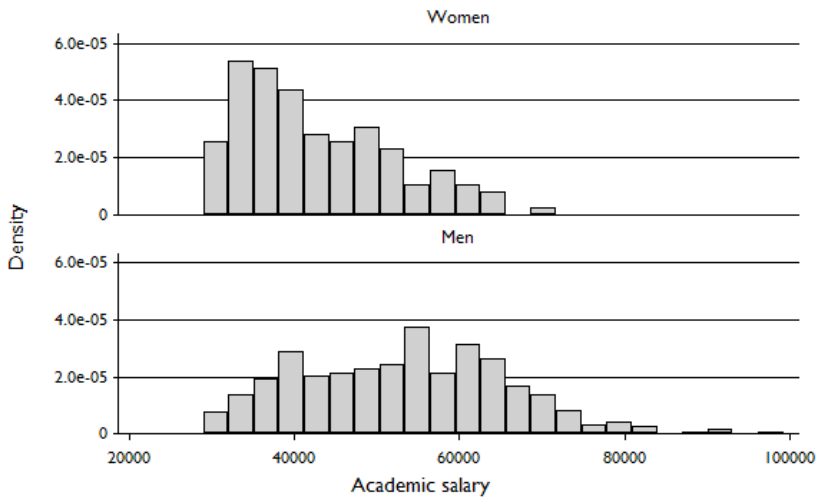
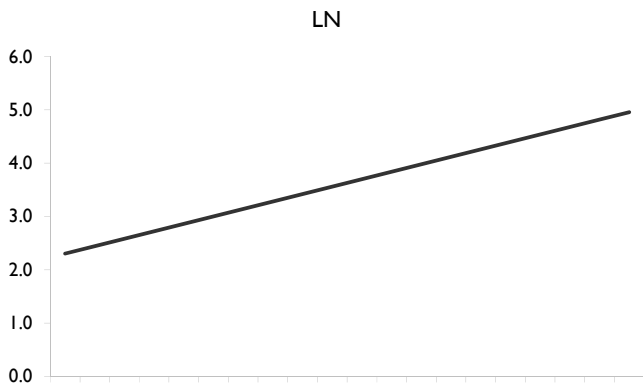
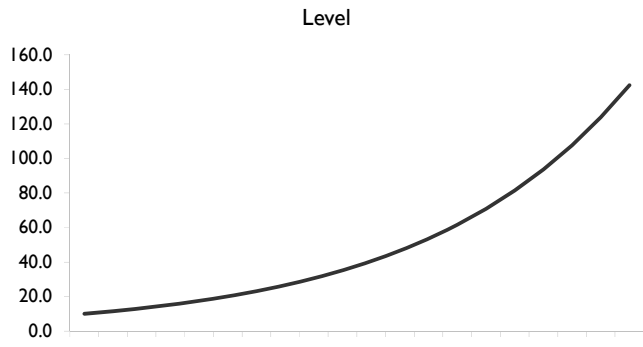
$$Y_i = \beta_0 + \beta_1 (1/X_{1i}) + \beta_2 X_{2i} + \varepsilon_i$$

Be sure to appropriately interpret the marginal effects. Elasticities, percentage changes etc.

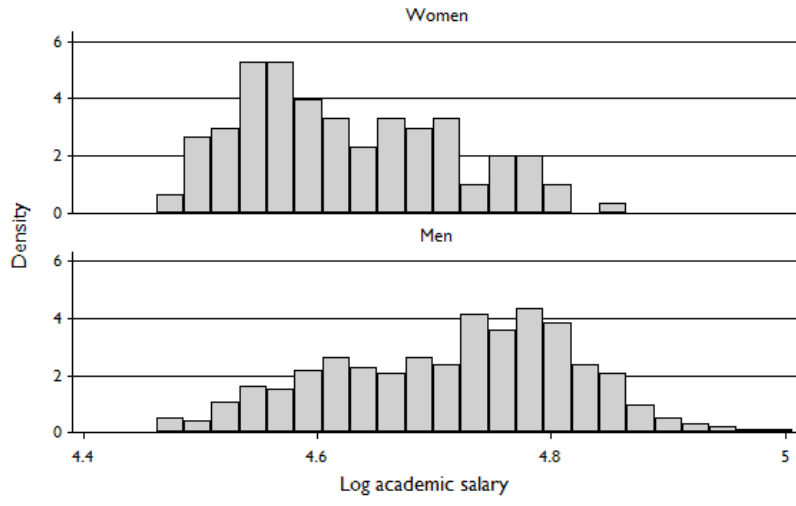
Never take the log of a dummy variable. Almost always take the log of a dollar value.

Problems with incorrect functional form.

Some pictures of alternative forms.



Graphs by male



Graphs by male

Rquared are difficult to compare when transformed

Incorrect functional forms

Estimate

Lecture 20 :November 7

Using dummy variables

Intercept dummy

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \varepsilon_i$$

Where:

Y is salary

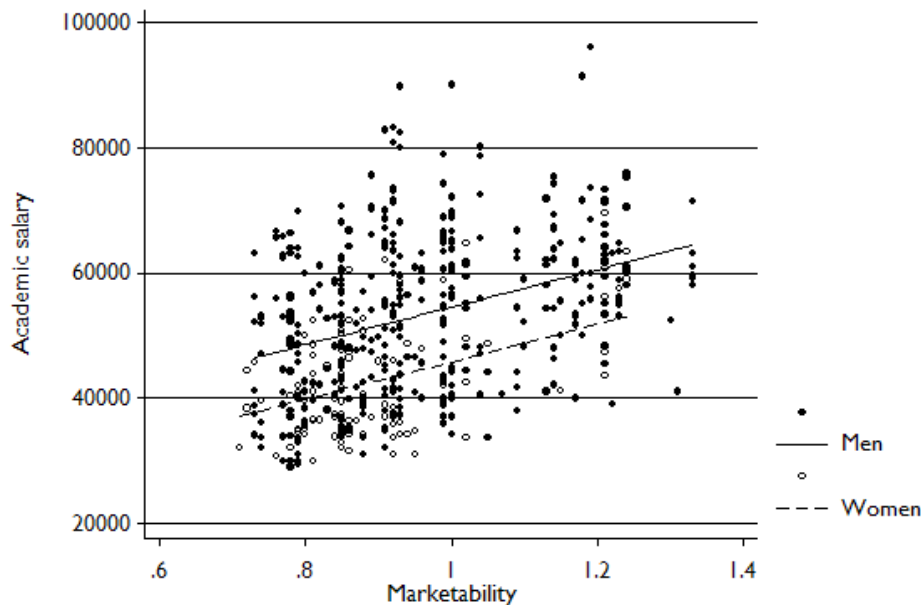
X1 is a dummy variable for male x2=1 for male, 0 for female.

X2 is marketability

. regress salary male marketc

Source	SS	df	MS			
Model	2.0711e+10	2	1.0356e+10	Number of obs =	514	
Residual	6.1676e+10	511	120696838	F(2, 511) =	85.80	
Total	8.2387e+10	513	160599133	Prob > F =	0.0000	
				R-squared =	0.2514	
				Adj R-squared =	0.2485	
				Root MSE =	10986	

salary	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
male	8708.423	1139.411	7.64	0.000	6469.917	10946.93
marketc	29972.6	3301.766	9.08	0.000	23485.89	36459.3
_cons	44324.09	983.3533	45.07	0.000	42392.17	46256



As a follow up from the previous section, I re-run the regression using the log of salary as the dependent variable. Notice a few things, the R-squared is different, but remember that should not be used to decide on models as the dependent variable has a different total sum of squares. Do notice that the coefficient on male is quantitatively different. Now its interpretation is the effect of being male not on salary, but the log of salary, or the percentage change. So being male means a 7.6% increase in salary relative to females holding market constant, but not other excluded/omitted variables.

. regress lsalary male marketc

Source	SS	df	MS			
Model	1.5890749	2	.79453745	Number of obs =	514	
Residual	4.44763545	511	.008703788	F(2, 511) =	91.29	
Total	6.03671035	513	.011767467	Prob > F =	0.0000	
				R-squared =	0.2632	
				Adj R-squared =	0.2604	
				Root MSE =	.09329	

lsalary	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
male	.0762761	.0096758	7.88	0.000	.0572669	.0952853
marketc	.2625476	.0280384	9.36	0.000	.207463	.3176323
_cons	4.635698	.0083506	555.14	0.000	4.619292	4.652104

Slope dummy (interaction terms)

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{3i} + \beta_4 X_{4i} \varepsilon_i$$

Where:

Y is salary

X1 is a dummy variable for male x2=1 for male, 0 for female.

X2 is marketability

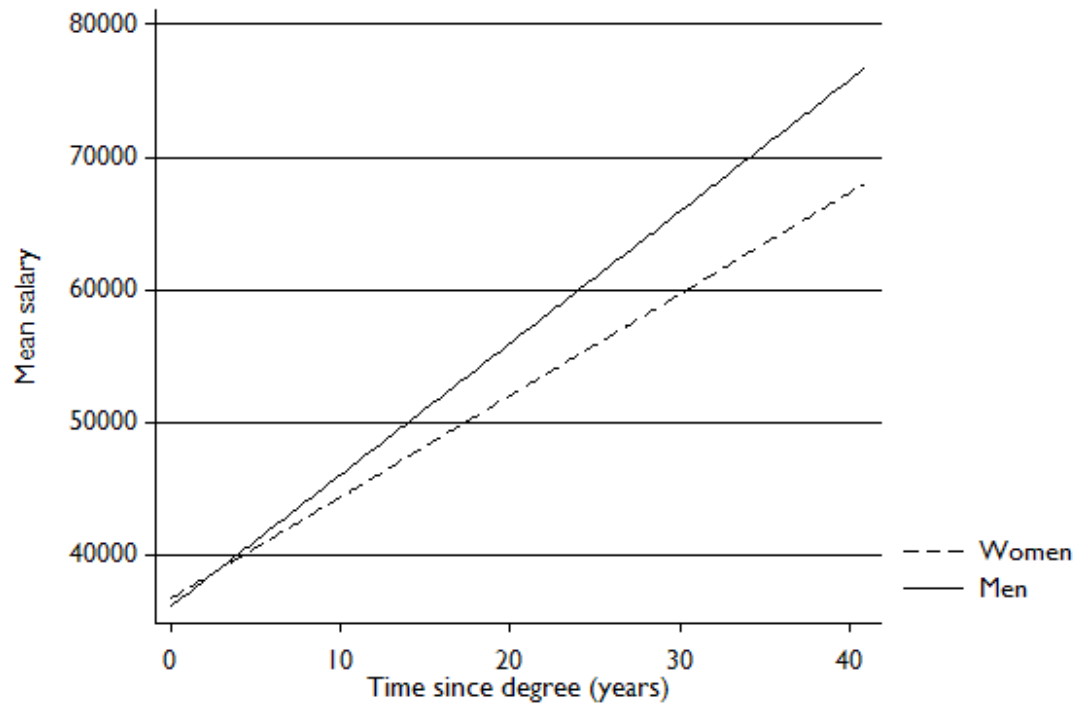
X3 is years to degree

X4 is m_years which is just X4=(X1*X3)

. regress salary male marketc yearsdg m_years

Source	SS	df	MS			
Model	5.6641e+10	4	1.4160e+10	Number of obs =	514	
Residual	2.5746e+10	509	50581607.4	F(4, 509) =	279.95	
Total	8.2387e+10	513	160599133	Prob > F =	0.0000	
				R-squared =	0.6875	
				Adj R-squared =	0.6850	
				Root MSE =	7112.1	

salary	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
male	-593.3088	1320.911	-0.45	0.654	-3188.418	2001.8
marketc	38436.65	2160.963	17.79	0.000	34191.14	42682.15
yearsdg	763.1896	83.4169	9.15	0.000	599.3057	927.0734
m_years	227.1532	91.99749	2.47	0.014	46.41164	407.8947
_cons	36773.64	1072.395	34.29	0.000	34666.78	38880.51



More uses for the F test.

Chow test

$$F = \frac{(RSS_T - (RSS_1 + RSS_2))/(k + 1)}{(RSS_1 + RSS_2)/(N_1 + N_2 - (2k + 2))}$$

where RSS_t is the residual sum of squares restricted equation and the others are from the individual unrestricted equations. It has an $F(k+1, n_1+n_2-2k-2)$ distribution.

Lecture 21: November 12
Multicollinearity

Lecture 22: November 14th

Remedies for MC

Do nothing

Drop redundant Variable

Transform variables

Increase sample size

Example